

Assessing the exceptionality of network motifs

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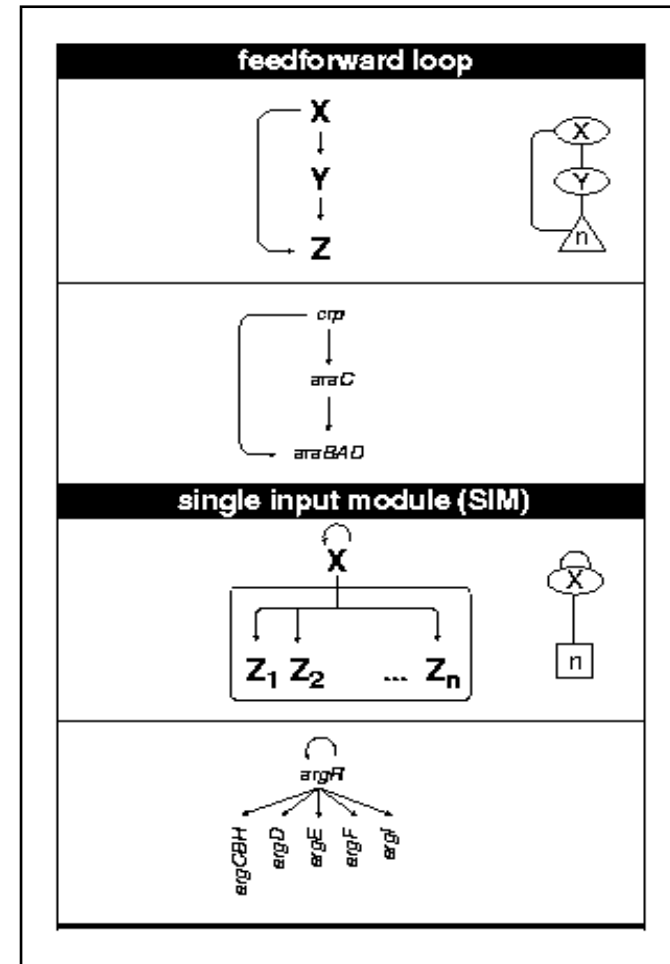
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Looking for local sub-structures in real-networks

- ▶ **Breaking-down complex networks into functional modules:**
 - patterns of interconnection,
 - **network motifs.**
- ▶ **Application in Biology:**
 - transcriptional regulatory modules
 - Example: feed-forward loop.
- ▶ **Exceptionality of a motif?**
 - when a given motif appears more frequently than **expected**.



From Shen-Orr et al. (2002)

How to assess the exceptionality of a motif

- ▶ Count the observed number $N_{\text{obs}}(\mathbf{m})$ of a given motif \mathbf{m} (out of our scope)
- ▶ Assess its significance with a p -value : need to calculate $\mathbb{P}\{N(\mathbf{m}) \geq N_{\text{obs}}(\mathbf{m})\}$
- ▶ Current strategy (Shen-Orr et al.):
 - use simulations to calculate $\mathbb{E}(N)$ and $\mathbb{V}(N)$ under a reference model
 - use a Z -score to calculate the p -value (implies a Gaussian approximation).

CONTRIBUTION

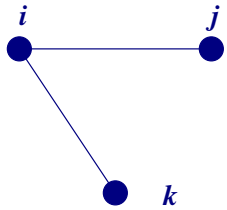
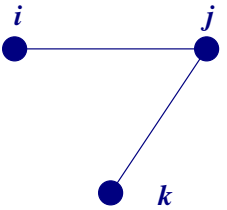
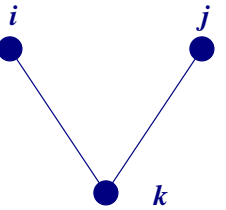
- 1 Give an analytic expression of the mean and the variance of the count,
- 2 Propose another distribution to better approximate the count distribution.

Definition of a motif in random graphs

- ▶ A random graph is defined by :
 - \mathcal{V} of fixed vertices with $|\mathcal{V}| = n$.
 - $\mathbf{X} = \{X_{ij}, (i, j) \in \mathcal{V}^2\}$ a set of random edges such that X_{ij} equals 1 if nodes i and j are connected, and 0 otherwise.
 - A distribution on X_{ij} . Example: the Erdős-Rényi model: $\mathbb{P}(X_{ij} = 1) = p$.
- ▶ **exchangeability hypothesis**: $\mathbb{P}(X_{ij})$ does not depend on (i, j) .
- ▶ \mathbf{m} stands for a motif of size k : connected subgraph with k vertices,
- ▶ It is defined by a fixed topology through its adjacency matrix also denoted by \mathbf{m} such that $\mathbf{m}_{uv} = 1$, if nodes u, v are connected in the motif

Example: the V motif

- ▶ 3 versions of the V motif at a **fixed** position $\alpha = (i, j, k)$.

| m | m' | m'' |
|--|---|--|
| $\begin{bmatrix} 0 & 1 & 1 \\ \cdot & 0 & 0 \\ \cdot & \cdot & 0 \end{bmatrix}$ | $\begin{bmatrix} 0 & 1 & 0 \\ \cdot & 0 & 1 \\ \cdot & \cdot & 0 \end{bmatrix}$ | $\begin{bmatrix} 0 & 0 & 1 \\ \cdot & 0 & 1 \\ \cdot & \cdot & 0 \end{bmatrix}$ |
|  |  |  |

Position and occurrence of a motif

- ▶ Let α be a possible position of \mathbf{m} . We consider that α is an ordered k -tuple with $i_1 < \dots < i_k$.
- ▶ We introduce the random indicator variable $Y_\alpha(\mathbf{m})$ which equals one if motif \mathbf{m} **occurs at position** α and 0 otherwise :

$$Y_\alpha(\mathbf{m}) = \prod_{1 \leq u < v \leq k} (X_{i_u i_v})^{m_{uv}} .$$

- ▶ Under the exchangeability assumption, the distribution of Y_α does not depend on α . Denoting $\mu(\mathbf{m})$ the **probability of occurrence** of motif \mathbf{m} , we have

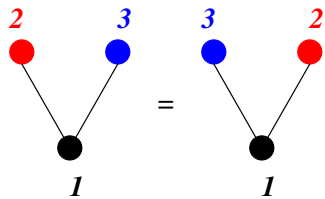
$$Y_\alpha(\mathbf{m}) \sim \mathcal{B}(\mu(\mathbf{m}))$$

- ▶ The number of occurrences of \mathbf{m} is then $N(\mathbf{m}) = \sum_{\alpha \in I_k} \sum? Y_\alpha(\mathbf{m})$.

Redundancy and Motif permutation

- ▶ For a given position, permutations of vertices of \mathbf{m} can lead to the same motif

$$\text{aut}(V) = \{\text{Id}, (2, 3)\}$$



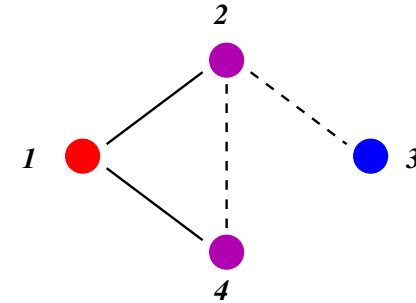
| | | | | |
|----------------------------|---|---|---|---|
| | | | | |
| $ \text{aut}(\mathbf{m}) $ | 2 | 6 | 6 | 8 |
| $\rho(\mathbf{m})$ | 3 | 1 | 4 | 3 |

- ▶ We define $\mathcal{R}(\mathbf{m})$, the set of non redundant permutations of \mathbf{m} , $\rho(\mathbf{m}) = |\mathcal{R}(\mathbf{m})|$.
- ▶ $\rho(\mathbf{m})$ is calculated with the $k!$ simultaneous permutations of the rows and columns of \mathbf{m} .
- ▶ The count of motif \mathbf{m} is: $N(\mathbf{m}) = \sum_{\alpha \in I_k} \sum_{\mathbf{m}' \in \mathcal{R}(\mathbf{m})} Y_{\alpha}(\mathbf{m}')$.
- ▶ We aim at calculating the mean of the count

$$\mathbb{E}N(\mathbf{m}) = |I_k| \times \sum_{\mathbf{m}' \in \mathcal{R}(\mathbf{m})} \mathbb{E}Y_{\alpha}(\mathbf{m}') = \binom{n}{k} \rho(\mathbf{m}) \mu(\mathbf{m}).$$

Calculating the variance is more intricate...

$$N^2(\mathbf{m}) = \sum_{\alpha, \beta \in I_k} \sum_{\mathbf{m}', \mathbf{m}'' \in \mathcal{R}(\mathbf{m})} Y_\alpha(\mathbf{m}') Y_\beta(\mathbf{m}'')$$



- ▶ \mathbf{m} : V motif
- ▶ \mathbf{m}' occurs at $\alpha = (1, 2, 4)$, \mathbf{m}'' occurs at $\beta = (2, 3, 4)$,
- ▶ In this case $\alpha \cap \beta = (2, 4)$
- ▶ The **super-motif** denoted by $\mathbf{m}' \underset{s}{\Omega} \mathbf{m}''$ is the union of two versions of \mathbf{m}
 - In this case, the super-motif is the so-called whisk graph motif
- ▶ We need to define:
 - the adjacency matrix of the super-motif $\mathbf{m}' \underset{s}{\Omega} \mathbf{m}''$
 - the non-redundant permutations of $\mathbf{m}' \underset{s}{\Omega} \mathbf{m}''$.

Adjacency matrix for super-motifs

- ▶ we break down \mathbf{m}' and \mathbf{m}'' such that:

$$\mathbf{m}' = \left(\begin{array}{c|c} \mathbf{m}'_{11} & \mathbf{m}'_{12} \\ \hline (k-s) \times (k-s) & (k-s) \times s \\ \hline \mathbf{m}'_{21} & \mathbf{m}'_{22} \\ \hline s \times (k-s) & s \times s \end{array} \right), \quad \mathbf{m}'' = \left(\begin{array}{c|c} \mathbf{m}''_{11} & \mathbf{m}''_{12} \\ \hline s \times s & s \times (k-s) \\ \hline \mathbf{m}''_{21} & \mathbf{m}''_{22} \\ \hline (k-s) \times s & (k-s) \times (k-s) \end{array} \right),$$

where \mathbf{m}'_{22} and \mathbf{m}''_{11} correspond to vertices in $\alpha \cap \beta$,

- ▶ We set:

$$\mathbf{m}' \underset{s}{\Omega} \mathbf{m}'' = \left(\begin{array}{c|c|c} \mathbf{m}'_{11} & \mathbf{m}'_{12} & \mathbf{0} \\ \hline \mathbf{m}'_{21} & \max(\mathbf{m}'_{22}, \mathbf{m}''_{11}) & \mathbf{m}''_{12} \\ \hline \mathbf{0} & \mathbf{m}''_{21} & \mathbf{m}''_{22} \end{array} \right).$$

- ▶ The \max function in the central term indicates that for the s common vertices of α and β , all edges of \mathbf{m}'_{22} and \mathbf{m}''_{11} must be present. It is equivalent to the logical OR.

New formulation for the squared count

- ▶ Each term of the sum depends on s , the number of shared vertices between α and β
- ▶ If $s = 0$, Y_α and Y_β are independent and $\mathbb{E}[Y_\alpha(\mathbf{m})Y_\beta(\mathbf{m})] = \mathbb{E}Y_\alpha(\mathbf{m})\mathbb{E}Y_\beta(\mathbf{m})$
- ▶ $\forall s \geq 1$, $Y_\alpha(\mathbf{m}')Y_\beta(\mathbf{m}'') = Y_{\alpha \cup \beta}(\mathbf{m}' \Omega_s \mathbf{m}'')$.
- ▶ The squared count can be rewritten as:

$$N^2(\mathbf{m}) = \sum_{s=0}^k \sum_{\substack{\alpha, \beta \in I_k : \\ |\alpha \cap \beta| = s}} \sum_{\mathbf{m}', \mathbf{m}'' \in \mathcal{R}(\mathbf{m})} Y_{\alpha \cup \beta}(\mathbf{m}' \Omega_s \mathbf{m}''),$$

- ▶ The expectation of the squared count is:

$$\mathbb{E}N^2(\mathbf{m}) = C_1(n, k) \left[\sum_{\mathbf{m}' \in \mathcal{R}(\mathbf{m})} \mu(\mathbf{m}') \right]^2 + \sum_{s=1}^k C_2(n, k, s) \sum_{\mathbf{m}', \mathbf{m}'' \in \mathcal{R}(\mathbf{m})} \mu(\mathbf{m}' \Omega_s \mathbf{m}'').$$

Calculating $\mu(\mathbf{m})$ in ER and ER mixture models

- ▶ $\mathbb{E}(N)$ and $\mathbb{V}(N)$ depend on $\mu(\mathbf{m})$ which depend on $\mathbb{P}\{\mathbf{X}\}$ (exchangeable)
- ▶ In the Erdős-Rényi model with parameter π , (m_{++} the number of edges in \mathbf{m}):

$$\mu_{\text{ER}}(\mathbf{m}) = \pi^{m_{++}}$$

- ▶ ERMG is an alternative model. We suppose that nodes are spread among Q hidden classes with proportion $\alpha_1, \dots, \alpha_Q$.

→ We denote by Z_i s the independent random variables which equal q if node i belongs to class q , then $X_{ij} | \{Z_i = q, Z_j = \ell\} \sim \mathcal{B}(\pi_{q\ell})$.

→ Under ERMG, we have:

$$\mu_{\text{ERMG}}(\mathbf{m}) = \sum_{c_1=1}^Q \dots \sum_{c_k=1}^Q \alpha_{c_1} \dots \alpha_{c_k} \prod_{1 \leq u < v \leq k} \pi_{c_u c_v}^{m_{uv}}.$$

Calculating $\mu(\mathbf{m})$ for Expected-Degree Distribution models

- ▶ EDD generates graphes whose degrees follow a given distribution,

$$\mathbb{P}\{X_{ij} = 1 | D_i D_j\} = \gamma D_i D_j$$

- ▶ "exchangeable" version of the Fixed Degree Distribution model (FDD),
- ▶ $\mu(\mathbf{m})$ can be calculated

$$\mu_{\text{EDD}}(\mathbf{m}) = \gamma^{m_{++}/2} \prod_{u=1}^k \mathbb{E} \left(D_{i_u}^{m_{u+}} \right).$$

- ▶ $\mu_{\text{EDD}}(\mathbf{m})$ only depends on the product of some moments of the expected degree D .

Comparison of theoretical moments on PPI networks

| Ecoli | N_{obs} | $\mathbb{E}N(\mathbf{m})$ | | | $\widehat{\mathbb{E}N(\mathbf{m})}$ |
|------------|------------------|---------------------------|--------------|--------------|-------------------------------------|
| | | ER | EDD | ERMG | FDD |
| V | 248,093 | 52,744.70 | 99,126.40 | 243,846.93 | 248,093 |
| triangle | 11,368 | 72.47 | 2,197.38 | 10,221.17 | 3,579.49 |
| chain | 9,557,956 | 399,151.00 | 2,339,200.00 | 9,555,414.55 | 5,950,903.40 |
| star | 6,425,495 | 133,050.00 | 1,537,740.00 | 5,772,005.15 | 6,425,495 |
| square | 487,408 | 411.31 | 38,890.60 | 417,190.55 | 76,467.39 |
| whisker | 2,154,048 | 1,645.22 | 306,789.00 | 1,929,516.68 | 547,802.44 |
| halfclique | 273,621 | 3.39 | 20,117.90 | 204,093.45 | 18,422.25 |
| clique | 14,882 | 0.00 | 867.24 | 8,904.75 | 317.27 |

First remark: the choice of the model has a strong influence on the first two moments.

This influence **depends on the topology of the motif.**

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- The expected count of **V** and **star** under **ER** and **EDD** are far from the observed count.
- Due to observed nodes with high degree which generate lots of occurrences of those motifs.

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| clique | 14,882 | 0.00 | 867.24 | 8,904.75 | 317.27 |

- The expected count under **ERMG** for **triangle**, **halfclique** and **clique** are close to the **observed** count
- Those motifs are linked to local clustering trends which are well captures by ERMG.

Comparison of theoretical moments on PPI networks

| Ecoli | N_{obs} | $\sqrt{\widehat{VN}(\mathbf{m})}$ | | | $\sqrt{\widehat{VN}(\mathbf{m})}$ |
|--------------|------------------|-----------------------------------|-----------|------------|-----------------------------------|
| | | ER | EDD-E | ERMG | FDD |
| V | 248093 | 1281.87 | 20851.70 | 51676.68 | 0 |
| triangle | 11368 | 8.90 | 797.30 | 3041.98 | 68.58 |
| chain | 9557956 | 14743.70 | 774109.00 | 3019630.93 | 67739.86 |
| star | 6425495 | 5089.62 | 484152.00 | 1672086.51 | 0 |
| square | 487408 | 29.14 | 19122.60 | 170502.21 | 1117.56 |
| whisker | 2154048 | 214.52 | 145764.00 | 739836.65 | 15593.00 |
| halfclique | 273621 | 2.04 | 12876.60 | 94018.80 | 891.99 |
| clique | 14882 | 0.05 | 707.94 | 4660.71 | 32.96 |

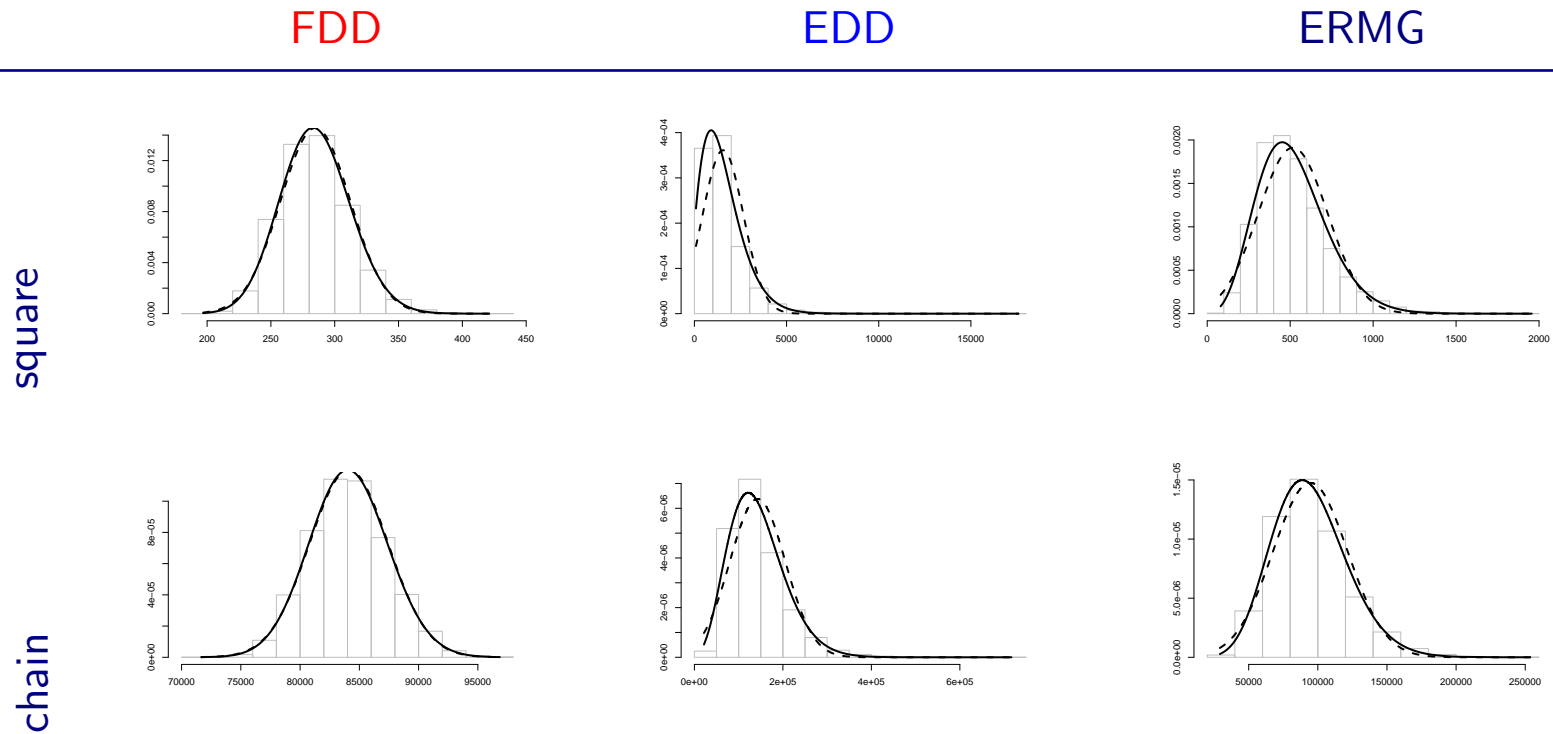
- When using the Fixed Degree Distribution model, the variance is systematically smaller
- Extreme case for the **V** and **star** motifs for which the degree exactly defines the number of occurrences

Compound Poisson approximation

- ▶ Exceptionality is assessed with $\mathbb{P}\{N(\mathbf{m}) \geq N_{\text{obs}}(\mathbf{m})\}$, where $N(\mathbf{m})$ the random number of occurrence of \mathbf{m} under the reference model.
- ▶ network motifs tend to overlap : clumps are present in the graph and C stand for the number of clumps (random)
- ▶ Denoting by S_i the size of clump i , we have $N(\mathbf{m}) = \sum_{i=1}^C S_i(\mathbf{m})$.
- ▶ If we make the hypothesis that $C \sim \mathcal{P}(\lambda)$, $N(\mathbf{m})$ is compound Poisson
 - We use the **Geometric-Poisson** distribution : we suppose that $S_i(\mathbf{m}) \approx \mathcal{G}(1 - a)$.
 - Then we approximate the distribution of $N(\mathbf{m}) \approx \mathcal{CP}(\lambda, a)$.
 - Parameters (λ, a) can be calculated according to $\mathbb{E}N(\mathbf{m})$ and $\mathbb{V}N(\mathbf{m})$

$$a = [\mathbb{E}N(\mathbf{m}) - \mathbb{V}N(\mathbf{m})]/[\mathbb{E}N(\mathbf{m}) + \mathbb{V}N(\mathbf{m})], \quad \lambda = (1 - a)\mathbb{E}N(\mathbf{m}).$$

Simulated count distributions



- The **shape** of the distribution highly depends on the model (whatever the motif)
- **FDD** generates **symmetrical** distributions (reflect the constraint of the model)
- **EDD** generates highly **skewed** distributions (diversity of visited configurations)

Conclusions for the simulation study

- ▶ Criteria used to assess the goodness of fit:
 - The Kolmogorov-Smirnoff distance between theoretical and empirical distributions
 - Empirical probabilities of exceeding the 0.999 quantile. It should be close to 0.001.
- ▶ The Geometric-Poisson approximation outperforms the Gaussian approximations for both criteria in all cases.
- ▶ The 0.999 quantile is underestimated by the Gaussian approximation:
 - the Gaussian approximation can lead to false positive results
- ▶ The KS distance is high for both approximations in some cases, especially for frequent and highly self overlapping motifs.
- ▶ However the clumps size distribution is not geometric...

Exceptional motifs in PPI networks

| Hpylo | N_{obs} | FDD-Pv \mathcal{P}_A | EDD-Pv \mathcal{P}_A | ERMG-Pv \mathcal{P}_A |
|------------|------------------|------------------------|------------------------|-------------------------|
| V | 14113 | - | 4.13e-01 | 4.06e-01 |
| triangle | 75 | 4.36e-03 | 9.06e-01 | 3.31e-01 |
| chain | 98697 | 1.22e-05 | 7.42e-01 | 4.12e-01 |
| star | 112490 | - | 3.65e-01 | 2.34e-01 |
| square | 1058 | 1.80e-52 | 6.15e-01 | 1.33e-02 |
| whisker | 3535 | 1.11e-02 | 8.58e-01 | 2.63e-01 |
| halfclique | 79 | 2.54e-05 | 7.51e-01 | 3.11e-02 |
| clique | 0 | 1.00e-00 | 1.00e-00 | 8.50e-01 |

- Using the FDD model leads to very drastic results (constant accross examples)
- When everything is exceptional the model should be questioned !
- For ERMG, 2 motifs are exceptional in the PPI network of H. Pylori

Exceptional motifs in PPI networks

| Ecoli | N_{obs} | FDD-Pv \mathcal{P}_A | EDD-Pv \mathcal{P}_A | ERMG-Pv \mathcal{P}_A |
|--------------|------------------|------------------------|------------------------|-------------------------|
| V | 248093 | - | 1.24e-08 | 4.46e-01 |
| triangle | 11368 | 0.00e+00 | 7.02e-13 | 3.30e-01 |
| chain | 9557956 | 0.00e+00 | 2.33e-10 | 4.68e-01 |
| star | 6425495 | - | 1.14e-11 | 3.26e-01 |
| square | 487408 | 0.00e+00 | 3.48e-23 | 3.10e-01 |
| whisker | 2154048 | 1.03e-265 | 1.15e-12 | 3.49e-01 |
| halfclique | 273621 | 1.24e-115 | 1.09e-17 | 2.14e-01 |
| clique | 14882 | 2.61e-41 | 3.30e-15 | 1.09e-01 |

- The behavior of the EDD model is not satisfactory
- Motifs are either all exceptional or all non exceptional
- May be linked to a variable quality of fit of the model to the data.

Exceptional motifs in PPI networks

| Scere | N_{obs} | FDD-Pv \mathcal{P}_A | EDD-Pv \mathcal{P}_A | ERMG-Pv \mathcal{P}_A |
|------------|------------------|------------------------|------------------------|-------------------------|
| V | 436131 | - | 6.21e-33 | 1.44e-01 |
| triangle | 10567 | 1.31e-128 | 1.13e-22 | 1.21e-06 |
| chain | 7530597 | 8.44e-99 | 8.61e-27 | 1.38e-01 |
| star | 12227236 | - | 3.11e-22 | 9.54e-03 |
| square | 165085 | 1.09e-322 | 3.19e-22 | 2.73e-02 |
| whisker | 993733 | 1.64e-65 | 9.33e-22 | 8.90e-04 |
| halfclique | 116667 | 1.71e-33 | 1.28e-18 | 7.22e-04 |
| clique | 8601 | 1.54e-10 | 3.19e-16 | 5.25e-06 |

- ERMG could be an alternative : Pvalues are moderate
- $\mu_{\text{ERMG}}(\mathbf{m})$ depends on the number of groups
- Model averaging to stabilize the procedure

Conclusions & future directions

- ▶ We propose a method to assess the exceptionality of network motifs.
- ▶ The method to calculate the moments of the count is general and can be applied to any random graph model with exchangeable distribution
- ▶ The Geometric-Poisson approximation for the count distribution works well on simulated data.
- ▶ Directions: how to assess the distribution of the clump size. Is there a general method or does it depend on each motif ?